

# ON THE INDEPENDENCE IN THE LIMIT OF SUMS DEPENDING ON THE SAME SEQUENCE OF INDEPENDENT RANDOM VARIABLES

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## Introduction

Let  $\xi_t$  be a stochastic process with independent increments. Suppose that  $\xi_t$  is integer-valued and its sample functions are continuous to the left and have a finite number of discontinuities with probability 1. It can be proved (see [3], Theorem 6) that if  $\nu_k$  is the number of discontinuities of  $\xi_t$  of magnitude  $k$  in the time interval  $I = [a, b]$ , then the random variables  $\nu_k$  ( $k = \pm 1, \pm 2, \dots$ ) are independent.<sup>1</sup>

This assertion implies, for example, that a homogeneous composed Poisson process  $\xi_t$  may be considered as a superposition of independent ordinary Poisson processes, i.e. can be represented in the form

$$\xi_t = \sum_{k=1}^{\infty} k \xi_t^{(k)},$$

where  $\xi_t^{(k)}$  is an ordinary homogeneous Poisson process, and the processes  $\xi_t^{(k)}$  are independent (see [4]). For a more general form of this statement see [3].

In § 1 of the present paper we prove a general theorem on the asymptotic independence of certain sums of random variables.

In § 2 deals with the application of our independence theorem leading to a theorem somewhat stronger than that formulated above. Further applications will be given in a forthcoming paper<sup>2</sup> of the first named author.

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<sup>1</sup>In [3] the above theorem is formulated more generally.

<sup>2</sup>On stochastic set functions. II, *Acta Math. Acad. Sci. Hung.*, **8** (1957), 337–374.

## § 1. The independence theorem

We start from a double sequence of random variables

$$\xi_{n1}, \xi_{n2}, \dots, \xi_{nk_n} \quad (n = 1, 2, \dots)$$

and suppose always that  $\xi_{n1}, \xi_{n2}, \dots, \xi_{nk_n}$  are independent for every  $n = 1, 2, \dots$ . Let us consider  $r$  Borel measurable real functions  $f_1(x), f_2(x), \dots, f_r(x)$  for which the sets defined by  $f_k(x) \neq 0$  are disjoint, or expressed in another way, for which the following relations hold:

$$(1) \quad f_j(x)f_k(x) = 0 \quad \text{for } j \neq k \quad (j, k = 1, 2, \dots, r).$$

Let us denote by  $\varphi_{lk}^{(n)}(u)$  the characteristic function of the random variable  $f_l(\xi_{nk})$ , further let us put

$$\zeta_l^{(n)} = \sum_{k=1}^{k_n} f_l(\xi_{nk}) \quad (l = 1, 2, \dots, r; n = 1, 2, \dots).$$

In order to simplify the understanding of the phenomenon which is described by our theorem, we formulate it first for a special case.

**THEOREM 1.a** *Let us suppose that the following conditions hold:*

- a) *The real, Borel measurable functions  $f_l(x)$  ( $l = 1, 2, \dots, r$ ) are integer-valued and satisfy (1).*
- b) *For every  $l$  ( $1 \leq l \leq r$ ) the random variables*

$$f_l(\xi_{n1}), f_l(\xi_{n2}), \dots, f_l(\xi_{nk_n})$$

*are infinitesimal, i.e.*

$$\lim_{n \rightarrow \infty} \sup_{1 \leq k \leq k_n} \mathbb{P}(f_l(\xi_{nk}) \neq 0) = 0.$$

- c.) *For every  $l$  ( $1 \leq l \leq r$ ) the limiting distribution of the random variables  $\zeta_l^{(n)}$  exists:*

$$(2) \quad F_i(x_i) = \lim_{n \rightarrow \infty} \mathbb{P}(\zeta_i^{(n)} < x_i) \quad (1 \leq i \leq r),$$

*at every point of continuity of  $F_i(x)$ .*

*Under these conditions the random variables  $\zeta_1^{(n)}, \zeta_2^{(n)}, \dots, \zeta_r^{(n)}$  are asymptotically independent, i.e.*

$$(3) \quad \lim_{n \rightarrow \infty} \mathbb{P}(\zeta_1^{(n)} < x_1, \zeta_2^{(n)} < x_2, \dots, \zeta_r^{(n)} < x_r) = F_1(x_1)F_2(x_2) \dots F_r(x_r)$$

*if  $x_i$  is a continuity point of  $F_i(x)$  ( $i = 1, 2, \dots, r$ ).*

PROOF. Let us consider the characteristic function of the joint distribution of the random variables  $\zeta_l^{(n)}$  ( $l = 1, 2, \dots, r$ ). Taking the relation (1) into account it can easily be seen by comparing the coefficients on both sides that the  $r$ -dimensional characteristic function of  $\zeta_1^{(n)}, \dots, \zeta_r^{(n)}$  is the following:

$$(4) \quad \sum_{j_1, j_2, \dots, j_r} \mathbb{P}(\zeta_1^{(n)} = j_1, \dots, \zeta_r^{(n)} = j_r) e^{i(u_1 j_1 + \dots + u_r j_r)} \\ = \prod_{k=1}^{k_n} \left\{ 1 + \sum_s \mathbb{P}(f_1(\xi_{nk}) = s) (e^{i u_1 s} - 1) + \dots + \sum_s \mathbb{P}(f_r(\xi_{nk}) = s) (e^{i u_r s} - 1) \right\}.$$

It follows from (4) that (denoting by  $\mathbb{M}(\chi)$  the expectation of  $\chi$ )

$$(5) \quad \mathbb{M}(e^{i(\zeta_1^{(n)} u_1 + \dots + \zeta_r^{(n)} u_r)}) = \prod_{k=1}^{k_n} \left\{ 1 + \left( \varphi_{1k}^{(n)}(u_1) - 1 \right) + \dots + \left( \varphi_{rk}^{(n)}(u_r) - 1 \right) \right\}.$$

Conditions a), b) and c) imply that the limits

$$(6) \quad \Phi_l(u_l) = \lim_{n \rightarrow \infty} \sum_{k=1}^{k_n} (\varphi_{lk}^{(n)}(u_l) - 1) \quad (l = 1, 2, \dots, r)$$

exist (see [1], § 24, Theorem 1) and  $e^{\Phi_l(u_l)}$  is the characteristic function of the limiting distribution  $F_l(x_l)$  ( $l = 1, 2, \dots, r$ ). Moreover, by Condition b) we have

$$(7) \quad \lim_{n \rightarrow \infty} \sum_{k=1}^{k_n} |1 - \varphi_{lk}^{(n)}(u_l)|^2 = 0 \quad (l = 1, 2, \dots, r).$$

According to (6) and (7) the sequence (5) converges to the  $r$ -dimensional characteristic function

$$e^{\Phi_1(u_1)} \dots e^{\Phi_r(u_r)}$$

and thus relation (3) holds.

A heuristic argument in favour of Theorem 1.a can be given as follows: Our suppositions a), b) and c) imply that in general only a small number of terms of the sum  $\zeta_l^{(n)} = \sum_{k=1}^{k_n} f_l(\xi_{nk})$  are different from 0 for each  $l$ . Supposition (1) ensures that the sums  $\zeta_l^{(n)}$  ( $l = 1, 2, \dots, r$ ) will always depend on disjoint subsets of the independent random variables  $\xi_{n1}, \xi_{n2}, \dots, \xi_{nk_n}$ ; of course, these sets are random, and therefore the sums  $\zeta_l^{(n)}$  are not independent, only almost independent. Nevertheless in the limit their dependence disappears.

The suppositions of Theorem 1.a may be replaced by a set of more special suppositions which, however, have the advantage that no supposition restricts at the same time the choice of the random variables  $\xi_{nk}$  and the choice of the functions  $f_l(x)$ , as there are two distinct groups of suppositions, further the convergence of the distribution of  $\zeta_l^{(n)}$  is not postulated, but is a consequence of the suppositions. This weaker form of Theorem 1.a is expressed by the following

COROLLARY Let  $\xi_{n1}, \xi_{n2}, \dots, \xi_{nk_n}$  denote a double sequence of independent non-negative integer-valued random variables which are infinitesimal, i.e.

$$\lim_{n \rightarrow \infty} \max_{1 \leq k \leq k_n} \mathbb{P}(\xi_{nk} \neq 0) = 0.$$

Let  $E_1, E_2, \dots, E_r$  denote disjoint subsets of the set of positive integers and let us suppose that  $f_l(k)$  ( $l = 1, 2, \dots, r; k = 0, 1, \dots$ ) are non-negative integer-valued functions such that  $f_l(0) = 0$  and  $f_l(k) = 0$  if  $k \notin E_l$ . Let us put  $p_{nks} = \mathbb{P}(\xi_{nk} = s)$ ,  $C_{ns} = \sum_{k=0}^{k_n} p_{nks}$  and suppose that there exists a convergent series of non-negative numbers  $\sum_{s=1}^{\infty} C_s$  such that

$$\lim_{n \rightarrow \infty} \sum_{s=1}^{\infty} |C_{ns} - C_s| = 0.$$

It follows that putting

$$\zeta_l^{(n)} = \sum_{k=1}^{k_n} f_l(\xi_{nk}) \quad (l = 1, 2, \dots, r; n = 1, 2, \dots)$$

we have

$$\lim_{n \rightarrow \infty} \mathbb{P}(\zeta_1^{(n)} < x_1, \zeta_2^{(n)} < x_2, \dots, \zeta_r^{(n)} < x_r) = F_1(x_1)F_2(x_2) \dots F_r(x_r),$$

where the distribution function  $F_k(x)$  has the generating function

$$\exp \sum_{s=1}^{\infty} C_s (z^{f_k(s)} - 1).$$

To prove that this Corollary really follows from Theorem 1.a, we have to apply Theorem 3 of the paper [2].

Now we turn to the general case in which the first part of Condition a) of Theorem 1.a is dropped. Our statement is expressed by

THEOREM 1.b *Let us suppose that the following conditions hold:*

a) *The Borel measurable real functions  $f_l(x)$  ( $l = 1, 2, \dots, r$ ) satisfy (1).*

b) *For every  $l$  ( $1 \leq l \leq r$ )*

$$\lim_{n \rightarrow \infty} \sum_{k=1}^{k_n} |\varphi_{lk}^{(n)}(u_l) - 1|^2 = 0.^3$$

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<sup>3</sup>It can be seen that if Conditions c) and d) hold, then Condition b) holds also if for some  $\tau > 0$

$$\lim_{n \rightarrow \infty} \sum_{k=1}^{k_n} |\alpha_{lk}^{(n)}|^2 = 0,$$

where

$$\alpha_{lk}^{(n)} = \int_{|x| < \tau} x dF_{lk}^{(n)}(x), \quad F_{lk}^{(n)}(x) = \mathbb{P}(f_l(\xi_{nk}) < x).$$

c) For every  $l$  ( $1 \leq l \leq r$ ) the random variables

$$f_l(\xi_{n1}), f_l(\xi_{n2}), \dots, f_l(\xi_{nk_n})$$

are infinitesimal, i.e. for every  $\varepsilon > 0$

$$\lim_{n \rightarrow \infty} \sup_{1 \leq k \leq k_n} \mathbb{P}(|f_l(\xi_{nk})| > \varepsilon) = 0.$$

d) For every  $l$  ( $1 \leq l \leq r$ ) the limiting distribution of the random variables  $\zeta_l^{(n)}$  exists.

Under these conditions the random variables  $\zeta_1^{(n)}, \zeta_2^{(n)}, \dots, \zeta_r^{(n)}$  are asymptotically independent, i.e. relation (2) holds.

PROOF. First we observe that (5) holds without the restriction that the  $f_l(x)$  are integer-valued. This can be shown as follows: By virtue of the independence of the variables  $\xi_{nk}$  we obtain

$$(8) \quad \mathbb{M} \left( e^{i \sum_{l=1}^r u_l \zeta_l^{(n)}} \right) = \prod_{k=1}^{k_n} \mathbb{M} \left( e^{i \sum_{l=1}^r u_l f_l(\xi_{nk})} \right).$$

Let  $A_{lk}^{(n)}$  denote the event consisting in that  $f_l(\xi_{nk}) \neq 0$ . Then we have<sup>4</sup>

$$(9) \quad \mathbb{M} \left( e^{i \sum_{l=1}^r u_l f_l(\xi_{nk})} \right) = \sum_{\nu=1}^r \left( \mathbb{M} \left( e^{i \sum_{l=1}^r u_l f_l(\xi_{nk})} \mid A_{\nu k}^{(n)} \right) - 1 \right) \mathbb{P}(A_{\nu k}^{(n)}) + 1.$$

As the event  $A_{\nu k}^{(n)}$  implies  $f_l(\xi_{nk}) = 0$  for  $l \neq \nu$ , we have

$$(10) \quad \mathbb{M} \left( e^{i \sum_{l=1}^r u_l f_l(\xi_{nk})} \mid A_{\nu k}^{(n)} \right) = \mathbb{M}(e^{iu_\nu f_\nu(\xi_{nk})} \mid A_{\nu k}^{(n)}).$$

On the other hand,

$$(11) \quad [\mathbb{M}(e^{iu_\nu f_\nu(\xi_{nk})} \mid A_{\nu k}^{(n)}) - 1] \mathbb{P}(A_{\nu k}^{(n)}) = \varphi_{\nu k}^{(n)}(u_\nu) - 1.$$

Thus (5) follows from (8)–(11).

Condition d) implies the existence of

$$(12) \quad \Psi_l(u_l) = \lim_{n \rightarrow \infty} \prod_{k=1}^{k_n} \varphi_{lk}^{(n)}(u_l) \quad (l = 1, 2, \dots, r).$$

As  $\Psi_l(u_l)$  is the characteristic function of an infinitely divisible distribution (see [1], § 24, Theorem 2), we have

$$\Psi_l(u_l) \neq 0 \quad (l = 1, 2, \dots, r)$$

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<sup>4</sup> $\mathbb{M}(\eta \mid A)$  denotes the conditional expectation of  $\eta$  under the condition  $A$ .

(see [1], § 17, Theorem 1). It follows hence and from (12) that if  $|\varphi_{lk}^{(n)}(u_l) - 1| \leq \frac{1}{2}$ , then

$$(13) \quad \left| \log \Psi_l(u_l) - \sum_{k=1}^{k_n} (\varphi_{lk}^{(n)}(u_l) - 1) \right| \leq \left| \log \Psi_l(u_l) - \prod_{k=1}^{k_n} \varphi_{lk}^{(n)}(u_l) \right| + \sum_{k=1}^{k_n} |\varphi_{lk}^{(n)}(u_l) - 1|^2 \quad (l = 1, 2, \dots, r).$$

The member on the right-hand side of (13) tends to 0, hence

$$(14) \quad \Phi_l(u_l) = \log \Psi_l(u_l) = \lim_{n \rightarrow \infty} \sum_{k=1}^{k_n} (\varphi_{lk}^{(n)}(u_l) - 1) \quad (l = 1, 2, \dots, r).$$

By (5) (14) and Condition b) it follows finally

$$\lim_{n \rightarrow \infty} \mathbb{M}(e^{i(\zeta_1^{(n)} u_1 + \dots + \zeta_r^{(n)} u_r)}) = \prod_{l=1}^r e^{\Phi_l(u_l)}.$$

Thus Theorem 1.b is proved. □

## § 2. Application to stochastic processes

In this § we consider a stochastic process with independent increments  $\xi_t$ . For the sake of simplicity we suppose that  $\xi_t$  is defined in the time interval  $[0, 1]$ . We suppose furthermore that the sample functions of  $\xi_t$  are continuous to the left for  $0 \leq t \leq 1$ , with probability 1. Let  $\nu(I)$  denote the random variable giving the number of discontinuities of  $\xi_t$  of magnitudes  $h \in 1$ . We prove the following

**THEOREM 2** *If the process  $\xi_t$  is weakly continuous, i.e. for every  $\varepsilon > 0$*

$$(15) \quad \lim_{\Delta t \rightarrow 0} \mathbb{P}(|\xi_{t+\Delta t} - \xi_t| \geq \varepsilon) = 0$$

*uniformly in  $t$  and  $I_1, I_2, \dots, I_r$  are pairwise disjoint intervals with positive distances from the point 0, then the random variables*

$$\nu(I_1), \nu(I_2), \dots, \nu(I_r)$$

*are independent.*

**PROOF.** Let  $f_l(x)$  denote the characteristic function (in the sense of set theory) of the interval  $I_l$ . We define the random variables

$$(16) \quad \eta_{n,k+1} = \xi_{\frac{k+1}{n}} - \xi_{\frac{k}{n}} \quad (k = 0, 1, 2, \dots, n-1).$$

Obviously,

$$(17) \quad \mathbb{P} \left( \nu(I_l) = \lim_{n \rightarrow \infty} \sum_{k=1}^n f_l(\eta_{n,k}) \right) = 1,$$

hence Condition c) of Theorem 1.a is satisfied. Since

$$\mathbb{P}(f_l(\eta_{n,k+1}) \neq 0) \leq \mathbb{P} \left( \left| \xi_{\frac{k+1}{n}} - \xi_{\frac{k}{n}} \right| \geq \delta \right),$$

where  $\delta$  is the minimal distance of the intervals  $I_l$  from the point 0, the random variables

$$f_l(\eta_{n,1}), f_l(\eta_{n,2}), \dots, f_l(\eta_{n,n})$$

are infinitesimal for every  $l$ . As Condition a) is obviously satisfied, the relations (2) and (17) imply our assertion.  $\square$

If instead of the intervals  $I_1, I_2, \dots, I_r$  we choose pairwise disjoint Borel measurable sets with positive distances from the point 0, then Theorem 2 holds obviously without any change. By choosing for  $f_l(x)$  other functions, further results can be obtained this way. For related results see [5].

## References

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